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# 統計的に有意な相互作用探索

#### 杉山 麿人(国立情報学研究所)









## **Example: Itemset Mining**

#### SNPs (items)





- ID 1: 00110011100111001110 ID 2: 11001011110001010100
- ID 3: 10100011101011000001
- ID 4: 1101101111111010011
- Control ID 5: 00110001100011111000 ID 6: 01011011000011001010 ID 7: 1011001010000101000
  - ID8: 11001001010100010101

### **Example: Itemset Mining**



## **Example: Subgraph Mining**



#### Timeline



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#### **Recent Advances**

- Webb, G.I., Petitjean, F.: A Multiple Test Correction for Streams and Cascades of Statistical Hypothesis Tests, KDD2016
- Pellegrina, L., Vandin, F.: Efficient Mining of the Most Significant Patterns with Permutation Testing, KDD2018
- Pellegrina, L., Riondato, M., Vandin, F.: SPuManTE: Significant Pattern Mining with Unconditional Testing, KDD2019
- Tran, T.Q., Fukuchi, K., Akimoto, Y., Sakuma, J.: Statistically Significant Pattern Mining with Ordinal Utility, KDD2020

### Libraries

- CASMAP
  - Llinares-Lopez, et al.: CASMAP: Detection of statistically significant combinations of SNPs in association mapping, Bioinformatics (2019)
- MP-LAMP (for parallel computation)
  - Yoshizoe, K., Terada, A., Tsuda, K.: MP-LAMP: parallel detection of statistically significant multi-loci markers on cloud platforms, Bioinformatics (2018)

## **Key Challenges:**

- 1. How to assess the significance for a multiplicative interaction of variables?
- 2. How to perform multiple testing correction?
  - How to control the FWER (family-wise error rate), the probability to detect one or more false positives?
- 3. How to manage combinatorial explosion  $(2^d \text{ for } d \text{ variables})$  of the candidate space?

## **Problem Formulation**

- Define  $X_{\mathcal{F}}$  as binary random variable of joint occurrence for a feature combination  $\mathcal{F} = \{F_i\}_{i \in I}, I \subseteq \{1, ..., d\}$ 
  - $X_{\mathcal{F}} = 1$  if  $\mathcal{F}$  "occurs",  $X_{\mathcal{F}} = 0$  otherwise
- Let *Y* be an output binary variable

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  - $X_{\mathcal{F}} = 1$  if  $\mathcal{F}$  "occurs",  $X_{\mathcal{F}} = 0$  otherwise
- Let *Y* be an output binary variable
- **Task:** Test the null hypothesis  $X_{\mathcal{F}} \perp Y$  for all  $\mathcal{F} \in 2^{V}$ 
  - Testing statistical independence between  $X_{\mathcal{F}}$  and Y

#### **Fisher's Exact Test**



## **Multiple Testing Correction**

- In each test, [probability of having a false positive]  $\leq \alpha$
- If we repeat *m* tests,  $\alpha m$  patterns can be false positives
  - Too many if *m* is large! For example in itemset mining:
    - For 100000 items, #patterns =  $2^{100000}$
    - Set significance level  $\alpha = 0.01$
    - Number of false positives:  $0.01 \cdot 2^{100000} = 10^{30101}$
- The FWER should be controlled
  - Probability at least one  $\mathcal{F}$  is false positive

## **Controlling the FWER**

- FWER = Pr(FP > 0)
  - FP: Number of false positives
- **Objective**: Maximize FWER( $\delta$ ) subject to FWER( $\delta$ )  $\leq \alpha$ 
  - FWER( $\delta$ ): FWER at corrected significance level  $\delta$ 
    - Cannot be evaluated in closed form (simple but not easy!)
  - Bonferroni correction is popular:  $\delta^*_{Bon} = \alpha/m$

• We use Tarone's testability trick, which requires the minimum achievable *p*-value  $\psi(\mathcal{F})$  for  $\mathcal{F}$ 

$$\psi(\mathcal{F}) = {\binom{N_1}{S(\mathcal{F})}} / {\binom{N}{S(\mathcal{F})}}$$
 in Fisher's exact test

#### $\mathcal{F}_{1}, \mathcal{F}_{2}, \mathcal{F}_{3}, ..., \mathcal{F}_{m-1}, \mathcal{F}_{m}, \mathcal{F}_{m+1}, ..., \mathcal{F}_{2^{d}} \quad \left(\psi(\mathcal{F}_{i}) \leq \psi(\mathcal{F}_{i+1})\right)$

 $m \psi(\mathcal{F}_m) < \alpha \text{ and } (m+1)\psi(\mathcal{F}_{m+1}) \ge \alpha$ 

 $\mathcal{F}_1\,,\,\mathcal{F}_2\,,\,\mathcal{F}_3\,\,,...,\,\,\mathcal{F}_{m-1}\,,\,\mathcal{F}_m\,,\,\,\mathcal{F}_{m+1}\,\,,...,\,\,\mathcal{F}_{2^d} \quad \left(\psi(\mathcal{F}_i) \leq \psi(\mathcal{F}_{i+1})\right)$ 



## Tarone's Testability Trick with Apriori

• We use Tarone's testability trick, which requires the minimum achievable *p*-value  $\psi(\mathcal{F})$  for  $\mathcal{F}$ 

$$\psi(\mathcal{F}) = {\binom{N_1}{S(\mathcal{F})}} / {\binom{N}{S(\mathcal{F})}}$$
 in Fisher's exact test

- This method is particularly effective if the relationship "Smaller  $\eta(\mathcal{F}) \rightarrow \text{Larger } \psi(\mathcal{F})$ " holds
  - For each pattern  $\mathcal{F}$ ,

 $S(\mathcal{F})$ : Support (how many times  $\mathcal{F}$  occurs in a dataset)  $\eta(\mathcal{F}) = S(\mathcal{F})/N$ : Frequency











## **Power of Testability**



The PTC (Predictive Toxicology Challenge) dataset with 601 chemical compounds

#### **Summary**

- 1. Formulate significance test for each pattern
  - Fisher's exact test is standard, while there are more possibilities
- 2. Enumerate testable patterns via Tarone's testability + Apriori (DFS)
- 3. Test each testable pattern

## LAMP and WY light



#### FACS [Papaxanthos et al. 2016] for Covariates



- Case ID 1: 00110011100111001110 Europe 1
  - ID 2: 11001011110001010100 Europe 1
  - ID 3: 10100011101011000001 Asia 1
  - ID 4: 1101101111111010011 Asia 1
- Control ID 5: 00110001100011111000 Europe 0
  - ID 6: 01011011000011001010 Europe 0
  - ID 7: 101100101000001010000 Asia 0
  - ID 8: 11001001010010010101 Asia 0



#### FAIS [Llinares-López et al. 2015] for Intervals



#### **FastCHM for Intervals + Cov.**



#### *C-Tarone* [Sugiyama & Borgwardt, 2019]

• Find all feature interactions form continuous data

Input:		У				
Г	F1	F2	F3	F4	F5	Class

- ID1 -0.96 -3.03 3.38 2.57 -6.06 ... 0
- ID2 –1.80 4.45 –4.35 0.82 8.90 ... 1
- ID3 -3.29 1.39 -4.44 -0.77 2.78 ... 1
- ID4 -0.53 -1.96 -3.43 -4.42 -3.92 ... 0

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#### *C-Tarone* [Sugiyama & Borgwardt, 2019]

• Find all feature interactions form continuous data

Input:			X				у	
	F1	F2	F3	F4	F5		Class	Output:
ID1	-0.96	-3.03	3.38	2.57	-6.06	•••	0	{F1}, {F3},
ID2	-1.80	4.45	-4.35	0.82	8.90	•••	1	→ {F2, F5},
ID3	-3.29	1.39	-4.44	-0.77	2.78	•••	1	{F2, F5, F6},
ID4	-0.53	-1.96	-3.43	-4.42	-3.92	•••	0	
•			:				:	

## Use Copula Support [Tatti, 2013]



Prod. 0.00  
0.11  
0.00  
0.22  

$$\int \frac{\text{Sum / 4}}{\text{O.083}} = \Pr(X_{\{\text{F1,F2,F3}\}} = 1) = \eta(\{\text{F1,F2,F3}\})$$
  
0.23/3

## **Contingency Tables**

- For each pattern (variable combination), we construct two types of contingency tables
  - One is from the expected situation under null
  - The other is from the observed situation from data
- Significance is assessed by comparison of the two tables
  - Each table is represented as a four-dimensional vector

	<b>Expected</b> for <b></b>	$\mathbf{p}_{\mathrm{E}}  X_{\mathcal{F}} = 1$	$X_{\mathcal{F}} = 0$	Total
	Y = 1	$\eta(\mathcal{F})r_1$	$r_1 - \eta(\mathcal{F}) r_1$	r <sub>1</sub>
	Y = 0	$\eta(\mathcal{F})r_0$	$r_0 - \eta(\mathcal{F}) r_0$	r <sub>0</sub>
	Total	$\eta(\mathcal{F})$	$1 - \eta(\mathcal{F})$	1
Oh	convod for m	V _ 1	V _ 0	Total
UD	served for $p_0$	$\Lambda_{\mathcal{F}} \equiv 1$	$X_{\mathcal{F}} \equiv 0$	TOLAT
	Y = 1	$\eta(\mathcal{F}, Y = 1)$	$r_1 - \eta(\mathcal{F}, Y =$	= 1) r <sub>1</sub>
	Y = 0	$\eta(\mathcal{F}, Y = 0)$	$r_0 - \eta(\mathcal{F}, Y =$	$= 0) r_0$
	Total	$\eta(\mathcal{F})$	$1 - \eta(\mathcal{F})$	1
				25/3

## Significance Test

- The independence  $X_{\mathcal{F}} \perp Y$  is translated into:
  - $H_0: D_{\text{KL}}(\boldsymbol{p}_{\text{O}}, \boldsymbol{p}_{\text{E}}) = 0, \quad H_1: D_{\text{KL}}(\boldsymbol{p}_{\text{O}}, \boldsymbol{p}_{\text{E}}) \neq 0$ 
    - $\boldsymbol{p}_{\rm E}$  and  $\boldsymbol{p}_{\rm O}$  are vectorized contingency tables:  $\boldsymbol{p}_{\rm E} = \left(\eta(\mathcal{F})r_1, \eta(\mathcal{F})r_0, r_1 - \eta(\mathcal{F})r_1, r_0 - \eta(\mathcal{F})r_0\right)$  $\boldsymbol{p}_{\rm O} = \left(\eta(\mathcal{F}, Y=1), \eta(\mathcal{F}, Y=0), r_1 - \eta(\mathcal{F}, Y=1), r_0 - \eta(\mathcal{F}, Y=0)\right)$
- We apply G-test: the statistic  $\lambda = 2ND_{\rm KL}(\boldsymbol{p}_{\rm O}, \boldsymbol{p}_{\rm E})$  follows the  $\chi^2$ -distribution with the d.f. 1

## **KL Divergence Bound**

- **Theorem** (tight upper bound of KL divergence):  $D_{\rm KL}(\boldsymbol{p}, \boldsymbol{p}_{\rm F})$ 
  - $< a \log \frac{1}{b} + (b-a) \log \frac{b-a}{(1-a)b} + (1-b) \log \frac{1}{(1-a)}$ -  $p_{\rm E} = (ab, a(1-b), (1-a)b, (1-a)(1-b)),$  $p \in \{ p \in \mathcal{P} \mid p_1 + p_2 = a, p_1 + p_3 = b \}$
- The *p*-value for this upper bound is the minimum achievable *p*-value

#### **Exp. Results on Synthetic Data**



#### Exp. Results on Synthetic Data



#### **Experimental Results on Real Data**



#### **Experimental Results on Real Data**



## Conclusion

- Significant pattern mining is introduced
  - Find significant interactions while controlling the FWER
  - pattern mining (data mining) + multiple testing correction (statistics)
- Key to solve the problem is Tarone's testability trick
  - This method can be used if the minimum achievable p-value  $\psi$  exists
  - If we have the relationship "Smaller  $\eta \rightarrow$  Larger  $\psi$ ", Apriori can be used to efficiently enumerate testable patterns 31/31